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Inequalities in Child Welfare Intervention Rates: the intersection of deprivation and identity

Paul Bywaters, Geraldine Brady, Tim Sparks and Elizabeth Bos

Abstract

Child welfare systems internationally exhibit very large inequalities in a variety of dimensions of practice, for example, in rates of child protection plans or registrations and out-of-home care. Previous research in the midlands region of England (Bywaters, 2013; Bywaters et al., 2014) has defined the concept of child welfare inequalities and detailed key aspects of the relationship between levels of neighbourhood deprivation and intervention rates. This paper reports further evidence from the study examining the intersection of deprivation with aspects of identity: gender, disability, ethnicity and age. Key findings include a decreasing gender gap and a decreasing proportion of children in need reported to be disabled as deprivation increases. The data challenge the perception that Black children are more likely than White to be in out-of-home care, a finding which only holds if the much higher level of deprivation amongst Black children is not taken in to account. Similarly, after controlling for deprivation and age, Asian children as found to be up to 6 times less likely to be in out-of-home care. The study requires replication and extension in order that observed inequalities are tested and explained. It provokes fundamental questions about whether higher rates mean that children's wellbeing and development are better secured: are Asian and Black families in deprived communities better able to support their children or are children's needs being neglected by service providers? Urgent ethical, research, policy and practice issues are raised about child welfare systems.

Keywords: child welfare, child protection, inequalities, deprivation, ethnicity, age, gender, disability

Inequalities in Child Welfare Intervention Rates: the intersection of deprivation and identity

Introduction

Child welfare systems internationally exhibit very large inequalities in a variety of respects (Bywaters, 2013; Bilson et al., 2013; Eckenrode et al., 2014; Bywaters et al., 2014). For example, child protection plan or registration rates varied between the four countries of the United Kingdom from 24.7 per 10000 children to 46.8 per 10000 in 2013¹. Child welfare inequalities, defined as ‘unequal chances, experiences and outcomes of child welfare that are systematically associated with social advantage/disadvantage’ (Bywaters, 2013, p.4) are seen in at least the following four aspects of practice.

- Inequalities in a parent’s or child’s chances of engagement with or intervention by child welfare services, reflecting diverse aspects of social position.
- Inequalities in the nature of child welfare interventions or provision for parents and/or children across different social groups or identities.
- Inequalities in childhood experiences and outcomes between different groups of children receiving welfare interventions and their counterparts in the wider population.
- Inequalities in outcomes as adults between different groups of children who received welfare interventions and those who did not.

These dimensions of child welfare inequalities, reflect similar markers of inequalities in health: inequalities in chances of illness and premature death, in access to treatment, in the quality, nature and outcomes of treatment. In each case the underlying issue, the reason why these are characterised as inequalities rather than just a lottery, is the systematic impact of social structural factors, such as social class, neighbourhood deprivation and ethnicity, on health. Similarly, child welfare intervention rates are markers of the impact of social structures on child wellbeing.

However, until recently, inequalities in child welfare have rarely been described in such terms, rather being talked about as variations, differences, disparities or disproportionalities. This way of describing inequalities may have contributed to masking structural relationships that lie at the roots of child welfare practice, relationships between social inequalities, rates of state intervention in family life and outcomes for children (Hearn, 2011). In turn, we suggest, this has contributed to a relative neglect in research, policy and practice of focus on fundamental causes (Phelan et al., 2010) of inequalities in access to child welfare services, in patterns of child welfare interventions and in outcomes, in favour of a ‘narrowed ...discourse ...that emphasises blame on individual behaviour’ (Stokes and Schmidt, 2011, p.1119). While ‘everybody in the business knows’ (Schorr, 1992, p.8) that access, interventions and outcomes are patterned by social inequalities, these factors are seen but not heard, universally familiar but rarely an explicit focus of policy and practice.

Nevertheless, in the last few years a focus on equity in access to, experience of and outcomes of child welfare structures and services has become more common (Tilbury and Thoburn, 2009). Much of this attention has been fuelled by an ongoing discussion of the relationship between socio-economic factors and racial identity in patterns of child welfare. For example, in the USA there has been a concern about the proportions of White, Black and Latino children in the child welfare system (Hill, 2007; Putnam-Hornstein et al., 2013; Wulczyn et al., 2013) and, in Australia, about rates of

intervention and outcomes for children from the majority and indigenous populations (Bilson et al., 2013). In Britain, too, racial ‘disproportionality’ has been the focus of some attention. Owen and Statham’s review (2009, p.1) stated that ‘It has long been known that children from black and mixed heritage backgrounds are over-represented among children who are looked after and Asian children tend to be under-represented.’ Usually, across national boundaries, studies have reported an over-representation of indigenous and Black or African-American children and an under-representation of other ethnic groups such as Latinos in the USA and ‘Asian’ children in the UK. Authors have continued to suggest a general ‘over-representation’ of ethnic minority children in care (Selwyn and Wijedesa, 2011) or in the child protection system (Ferguson, forthcoming).

Over the past ten years or so, researchers in the USA have begun to question these traditional certainties. Studies have shown that, when controlled for socio-economic status (SES), apparently higher rates are diminished or even reversed for Black US children by comparison with low SES White children (Putnam-Hornstein et al., 2013). Indeed, recent evidence suggests that ‘social disadvantage... is correlated with placement rates more for whites than for blacks’ (Wulczyn et al., 2013, p73). Further, when controlled for SES, rates for Latino children remain significantly lower than those for White children (Putnam-Hornstein et al., 2013, p. 44), suggesting real differences between ethnic groups in ‘maltreatment risk’ for equivalent levels of deprivation, that disparities/inequalities are not just the result of systemic or practice bias.

Such investigations are valuable for two main reasons: they shed light on alternative explanations for the variable rates of intervention found between children of different identities and in different locations and they provide a more nuanced and, therefore, more precise account of the working of the child welfare system. Hence, they create a firmer foundation for policy, planning, resource allocation, professional education and practice.

A key concept here is ‘intersectionality’. As a recent review of the relationship between poverty and ethnicity put it:

People’s experience is not shaped by one aspect of their identity alone but by a combination of elements. Gender, age, religion, disability, health, location and migration history can all be as important as ethnicity. They can change how ethnicity affects people’s self-perception and treatment by others. In addition, class, skills and qualifications, personal outlook and experience can change the meaning that such demographic characteristics have. (Barnard and Turner, 2011, p.4)

Studies of child welfare systems which focus on multiple dimensions of identity remain rare but to understand the complexities of the relationship between deprivation and differential rates of child welfare intervention, it is necessary to take into account aspects of identity. This means not only considering ethnic identity, but also other intersecting factors suggested by Barnard and Turner (2011) above. However, in England, data is routinely recorded by child welfare services on only some dimensions of identity: age, gender, disability and ethnicity, but not, for example, religion or sexual orientation, either of which may be of importance to the children themselves and to how they are treated (Fish, 2013). There is no internationally collected standard data set on child welfare systems.

This paper reports on issues of intersectionality as they inform the pattern and nature of child welfare inequalities in England. In 2013, research funded by the Nuffield Foundation, began to examine differential rates of children's services interventions between and within local authorities (LAs) in England. A previous paper, (Bywaters et al., 2014) reported on three key findings relating neighbourhood deprivation to rates of intervention.

- Very large inequalities in rates of children's services interventions were linked to deprivation. Children's chances of being a looked-after-child (LAC) or on a Child Protection Plan (CPP) varied greatly between neighbourhoods, as well as between LAs. Unequal rates were strongly statistically associated with measures of deprivation.
- A social gradient in child welfare interventions. Despite the large increase in rates associated with greater deprivation, children from families across the whole of society were LAC or on CPPs. Almost 40% of LAC or CPP came from outside the most disadvantaged 20% of neighbourhoods.
- An 'inverse intervention law' for child welfare. Overall a child's chances of a child welfare intervention increased significantly with deprivation but, comparing equivalent levels of neighbourhood deprivation, a child in a more affluent local authority overall was more likely to be on a CPP or to be LAC.

Here we report findings on the intersection between neighbourhood disadvantage and four aspects of identity: disability, gender, ethnicity and age. The paper addresses only the first of the four markers of child welfare inequalities: unequal rates of intervention.

Methods

Data was provided by 13 LAs in the English midlands, a mixture of urban boroughs and more rural counties. The participating LAs were responsible for nearly 1.2 million children aged 0-17, 10.5% of all children in England, 10.6% of all children on a CPP and 11.3% of all LAC on March 31st, 2012. Each LA reported on the age, gender, ethnic group and disability of all children in need (CIN), on a CPP and LAC on the given date. This data mirrored that routinely provided annually by LAs to produce national statistics. In addition, LAs identified the neighbourhood in which each child lived or, for LAC, of the address from which they became looked after. The neighbourhoods were those known as 'lower super output areas' (LSOAs), covering an average of 1500 residents, an element of the national structure of geographies on which official statistics are based.

Having amalgamated the data into LSOAs (n=3252), we analysed the relationships between rates of intervention and deprivation using population counts drawn from the 2011 Census and Index of Multiple Deprivation (IMD) scores for LSOAs published in 2010. The IMD is a broad measure of deprivation encompassing 7 key dimensions and 38 indicators, not solely a measure of income. The primary form of analysis involved grouping LSOAs into deprivation quintiles ranked in terms of the national IMD scores. In subsequent tables and charts showing results for quintiles 1 to 5, results for quintile 1 refer to all those neighbourhoods in the sample which were amongst the 20% least deprived neighbourhoods nationally. Quintile 5 refers to those neighbourhoods which were amongst the 20% most deprived neighbourhoods nationally. For much of this paper, data was analysed at middle layer super output area (MSOA) level because of small numbers in relevant cells once multiple variables are considered. MSOAs are groups of LSOAs with an average population of 7200. The study methods are described in detail in the earlier paper (Bywaters et al., 2014).

Findings

Deprivation and the Child Population

Central to this article is the thesis that understanding patterns of demand for and intervention by child welfare services requires examining the relationship between deprivation and aspects of identity. While national and local policies, cultures and practices play an important part in determining the support families receive from child welfare services or their sometimes unwelcome interventions, structural factors frame patterns of demand. Our findings suggest that two key factors – deprivation and ethnicity – play the largest role in determining inequalities in rates of intervention, while other factors are also important contributors.

The UK is an unequal society. OECD data for 2010 show that the UK came 19th amongst European nations for the proportion of children living in relative poverty (<http://www.oecd.org/social/family/database/CWBM>). In 2011, 17% of UK children were living below the poverty line. This is anticipated to grow to 24% (3.3 million children) before housing costs by 2020 (Brewer et al, 2011). Currently, multiple policies aimed at reducing state expenditure have worsened and will continue to worsen the circumstances of children (Reed, 2012). Discussing inequalities rather than just poverty is significant here. Children's circumstances are more accurately described as being on a continuum from the very poor to the very wealthy, rather than as a binary divide between those in poverty and those who are not.

The broad pattern nationally (Table 1) is that children are disproportionately found amongst relatively poor households. If distributed equally, each quintile of neighbourhoods should contain 20% of the population as a whole and of children but 23.7% of children were living in the most disadvantaged 20% of neighbourhoods at the 2011 Census. This pattern of excess deprivation in childhood was exaggerated in our midlands sample, which contained a disproportionate number of disadvantaged neighbourhoods. The combination of these two factors means that, in our sample, almost 2 in 5 children lived in the most disadvantaged fifth of neighbourhoods nationally. For very young children (aged 0 to 4) the over-representation was even greater: 4 times more young children lived in quintile 5 than in quintile 1.

There were also major differences in population patterns between LAs. In Birmingham and Sandwell over 60% of all children lived in the most deprived quintile, while in Warwickshire and Herefordshire this was true for less than 10% of children. Given that children in the 5th quintile were over 7 times more likely to be LAC and over 9 times more likely to be CPP compared to quintile 1, the significance of the distribution of children amongst more or less deprived neighbourhoods for demand for children's services is apparent.

Gender

Overall, in our sample, boys were a little more likely than girls to be the subject of child welfare interventions. The gender gap (excess proportion of boys over girls) was 7.2% for CIN, and 5.5% for LAC, but only 2.7% for CPP. These differences were spread remarkably evenly across age groups and CIN, CPP and LAC. The slightly greater preponderance of boys also applied to all five broad ethnic groups. Numbers of girls (n=54) exceeded boys (n=46) amongst children aged 16 and 17 on CPPs, but a handful of children could have reversed the pattern. There were also a very few more girls than

boys on CPPs in the most disadvantaged decile. Indeed, there was some evidence of a diminishing gap between boys and girls as deprivation increased (see Chart 1 for CIN) but this was a relatively slight relationship compared to other dimensions of identity, and only applied to CIN. For CIN, chi-square analysis of the number of boys and girls against quintile ($\chi^2(4)=11.10$, $P=0.025$) supports the case that the gender gap reduced as deprivation increased. No obvious explanation for this was available from this study.

Disability

The proportion of children in England who are disabled is a contested area, partly because of differing definitions and data sources. Read et al. (2010, p. 142) in the most authoritative recent UK study of the prevalence of disability in childhood, defined it as 'children with a limiting longstanding (12 month duration or longer) illness, disability or infirmity experiencing one or more significant difficulties or health problems (...including...) those who would have such difficulties or problems if they did not take medication/s.' They reported that 7.3% of children in the UK were disabled, 8.8% of boys and 5.8% of girls. However, the 2011 Census (Table DC3302EW) found that only 1.5% of children in England aged 0-15 were reported by parents to have a disability which limited their day-to-day activities 'a lot' with a further 2.2% having their activities limited 'a little'.

Mooney et al. (2008) found that local authorities were unable to 'provide a reliable figure for the number of disabled children ... due to the difficulties LAs had in identifying and counting disabled children and the different definitions of disability applied' (p.5). However, they estimated on the basis of a number of data sources that between 3% and 5.4% of all children in England were disabled.

The annual Children in Need census asks LAs to record whether a child is disabled or not using the following definition drawn from the Disability Discrimination Act 2005: 'a physical or mental impairment which has a substantial and long-term adverse effect on his ability to carry out normal day-to-day activities...the condition has lasted or is likely to last at least 12 months in order to be counted as a disability.' However, children who only 'receive mail outs (for example newsletters)' were excluded (Department for Education, 2012, p.7). This guidance was adopted for this study. Many, perhaps most, disabled children will not be receiving social care services and this is reflected in our finding that disabled children who were CIN constituted only 0.32% of all children in the participating LAs. Using Blackburn et al.'s (2010) data, less than 1 disabled child in 20 was designated a CIN.

Our findings, too, suggest very different practices between LAs in how disability was recorded. The proportion of CIN recorded as disabled varied from 6.7% to 23.6%. This does not seem plausible as a real difference in prevalence between LAs. Similar large differences between LAs were found for CPP and LAC.

The overall proportion of children recorded as disabled is lowest for children on CPPs. This partly reflects a wider issue: rates of recorded disability are much lower for children under 5 than for other ages (Table 2). Blackburn et al. (2010) found that the UK rate of disability for under-fives was under half that for older children and suggested a number of reasons including possible under-recording. Diagnoses of disability may take some time to be confirmed, with more cases emerging as children enter school. In our sample, almost 45% of all children on CPPs were under 5, compared to less than

30% for CIN and LAC. An alternative explanation is that the focus on child protection over-rides attention to issues of disability. It again seems implausible that children on CPPs are actually less likely to be disabled than CIN or LAC.

Overall, the proportion of Asian CIN recorded as disabled was almost twice (24%) that for the other ethnic groups (12.5%). A similar difference was found for CPP but amongst LAC the proportion of Asian children who were disabled was lower than for other ethnic groups. However, in both these latter cases numbers were very small ($n < 20$ for disabled Asian children).

As Table 3 (for CIN alone) and Chart 2 (for CIN, CPP and LAC) indicate, the intersection of disability and deprivation appears to show progressively lower proportions of disabled children as deprivation increases. Disabled children in more affluent areas were more likely to be CIN than disabled children in more deprived areas. Given the evidence that disabled children live in 'substantially more disadvantaged material circumstances than () non-disabled children' (Blackburn et al., 2010, p.9) this is the reverse of what would be expected. It is, perhaps, another example of Tudor Hart's (1971) inverse care law applying to child welfare services (Bywaters et al., 2014).

Although a similar but shallower gradient can be seen (Chart 2) for CPP and LAC, the relationship between intervention rates and deprivation is not statistically significant (for LAC: $F_{1,12} = 1.91$, $P = 0.192$), with numbers very small. The difference in gradient between CIN, CPP and LAC might be because parents of disabled children seek services which require CIN status but not for their child to be either on a CPP or a LAC, with parents from more affluent areas being more effective at securing services. This finding requires replication and explanation.

Ethnicity

Before presenting the findings on ethnicity, a number of caveats must be entered. For the purposes of statistics collected on CIN, CPP and LAC in England, the term 'Black' is used to mean 'Caribbean, African or any other Black background' and to exclude children of Mixed heritage such as 'White and Black Caribbean' or 'White and Black African' (Department for Education, 2012). The term 'Asian' refers to children of 'Indian, Pakistani, Bangladeshi or any other Asian background' and excludes children of mixed 'White and Asian background'. It can immediately be seen that these categories are problematic, as people of Mixed heritage may identify themselves as Black or Asian. It is also not known how consistently social workers or other social services staff collect this data.

A further important limitation is the amalgamation of ethnic groups into very broad categories. Such categories deny the diverse lived experience and identities of people in all ethnic groups and obscure the material circumstances of different sub-groups. For example, in Wolverhampton over 70% of the 'Asian' population in the 2011 census identified themselves as of Indian heritage and only 10% as Pakistani, while in Birmingham 23% identified as Indian and over 50% as Pakistani. Similarly, in Coventry over 70% of the population designated 'Black' identified themselves as of African heritage compared to less than 25% in Sandwell. These differences are important as rates of poverty and many other factors vary substantially (Barnard and Turner, 2011). Future research needs to drill down further into ethnic groupings.

Those limitations aside, the aggregate national data for 2012 (Table 4) shows a pattern similar to that reported by Owen and Statham (2009) for the period 2003 to 2006. That is, there were

substantially higher rates of CIN and LAC for Black children than for White but a much closer relationship for CPP, with children of Mixed heritage experiencing the highest rates. Significantly lower rates were reported for Asian children in all three forms of intervention.

The overall data for our midlands sample (Table 5) shows both similarities and differences with the national data. Although LAC rates for Black children were above those for White children, rates for both CPP and CIN were lower. Rates for children of Mixed heritage were the highest in each form of intervention while rates for Asian children were lowest.

However, examining this broad picture in more depth, a central issue is the distribution of the population of ethnic groups by deprivation. While all children were over-represented in the most disadvantaged quintile of neighbourhoods, there was a striking difference between White and other ethnic groups in this respect (Table 6). Over two thirds of Asian children and over three quarters of all Black children lived in the most disadvantaged 20% of neighbourhoods. By contrast little more than a quarter of all White children lived in these areas.

So, after controlling for deprivation by comparing ethnic groups within each deprivation quintile (Table 7), it is not Black children who were over-represented in rates of CIN, CPP or LAC but White children and those of Mixed heritage. Black rates in neighbourhoods in quintile 5, where over 76% of all Black children were living, were around half of those for White children for CIN and CPP and less than two thirds for LAC. It is only because so many Black children were living in the most deprived neighbourhoods where intervention rates were higher that it appears that they are over-represented in the system overall. Because previous studies of Black children's chances of being a LAC have not taken the levels of deprivation in the population into account, a misleading impression has been established over many years.

Previous research has detailed the under-representation of Asian children in child welfare interventions (Owen and Statham, 2009). Comparing Asian and White children at the same level of deprivation greatly increases the size of the disparity (Table 7). In quintile 5, where two thirds of Asian children lived, CPP rates for Asian children were three times lower than those of White children and LAC rates around six times lower. The ratios in quintile 4 were approximately 2.5 and 4.5, respectively. Explanations for such very large differences have been mooted but not yet demonstrated. The previously observed gap between White and Mixed heritage children is also reduced, although still present in the most disadvantaged fourth and fifth quintiles where over 70% of Mixed heritage children were living.

Chart 3 shows ethnic differences in LAC rates across the whole midlands sample (excluding the 'Other' ethnicity group because of low numbers). A number of features of our analysis can be seen here. First, this graphically represents the much higher rates amongst White and Mixed heritage children in the fourth and fifth quintiles, compared to Asian and Black children. Second, it can be seen that the gradient – a measure of the strength of the relationship between deprivation level and rates - is steeper for White and Mixed heritage children than for Black and Asian children. This finding requires further confirmation, especially as the numbers of Black and Asian children in quintiles 1 to 3 are small, only 57 children in all. However, it may mirror Wulczyn et al.'s conclusion that in the USA that 'social disadvantage ... is correlated with placement rates more for whites than blacks' (2013, p.73). Again, explanation is required. Third, the graph suggests that in neighbourhoods

of low deprivation, LAC rates for Black children were higher than those for White. This, too, requires replication as the numbers involved were very small (just 29 Black children).

The data also adds detail to the ‘inverse intervention law’ reported previously (Bywaters et al., 2014). While the inverse statistical relationship between overall IMD scores at the LA level and rates of intervention at any given level of deprivation holds true for White children, indeed the relationship is even stronger, it is not found for other ethnic groups, mirroring Wulczyn et al. (2013). This, like the shallower gradient, may reflect a less strong relationship between social disadvantage and welfare intervention rates amongst minority ethnic groups, for reasons as yet unknown.

Age

Age is a further dimension in unravelling inequalities in the child welfare system. Age is a significant variable in both CPP and LAC rates (Table 8, 9). As children age, their chance of being on a CPP fall, in all deprivation quintiles and overall. CPPs are much more common amongst children under 5 than when older and, at that age, CPP rates are close to LAC rates. It is the much greater drop in CPP rates as children age by comparison with LAC rates, which results in the lower overall CPP rate. For LAC, rates at a point in time are highest for the oldest children.

If you combine CPP and LAC rates – in our sample less than 2.5% of cases were duplicated – the overall rate at which the state intervenes most powerfully to safeguard children is quite similar at different ages after the early years but the balance shifts from CPP to LAC as the predominant form of intervention (Table 10). The nature of this interaction, for example, whether these are in a sense the same children at different ages or different children, is one that could be explored further and might have a bearing on arguments about early intervention. The patterns in which children move in and out of CPP or LAC status at different ages will be an important part of the emerging picture.

The proportion of children in different ethnic groups also varies with age. Tables 11 and 12 show how age, ethnicity and deprivation intersect to produce inequalities in intervention rates. Overall the CPP rate for White children was nearly twice that for Asian children. But for children under 5 in quintile 5, a White child was more than 4 times more likely to be on a CPP than an Asian child. Similarly the higher rate for White compared to Black children increases from a ratio of 1.2 overall, to 2.2 for all ages in Quintile 5 in which most Black and Asian children live, and to 2.6 for children under 5. Adding age to deprivation further increases the inequalities in rates between ethnic groups.

Similarly (Table 12), while the overall ratio of White to Asian LAC rates was 3.6, for the youngest children in quintile 5, the ratio increased to 6.6. For Black children, the slight excess of Black compared to White children overall, is much more than reversed for children under 10 in quintile 5, where the rate for White children was around twice that for Black children.

Discussion and Conclusions

Some of our findings add detail while confirming previous evidence, some overturn long held assumptions, others raise new questions about the relationships between deprivation and intervention rates. Overall, our data underline the growing case that analysing child welfare intervention rates without taking detailed account of the intersection between deprivation and identity leads to a misleading conception of current child welfare systems. As Franzen et al. (2008,

p.1057) put it, 'Beware of reports of over-representation of minorities in child welfare populations if the figures are presented without adjustments for socio-economic background.'

We would draw four key points from the data.

1. For most kinds of intervention studied there was a small excess of boys over girls and this was consistent across age and ethnic groups and types of intervention, with few exceptions. However, for CIN the gender gap reduced as deprivation increased.
2. Data recording on disability is unreliable. Unexplained differences between the proportions of disabled children amongst CIN, CPP or LAC, between LAs and between Asian and other ethnic groups need further examination. The proportion of CIN reported as disabled decreased as deprivation increased, contrary to expectations.
3. In this sample, ethnic minority children were very greatly over-represented in high deprivation neighbourhoods compared to White children. A more detailed examination of ethnic identity below the broad groupings would be valuable.
CIN, CPP and LAC rates for Black children in quintiles 4 and 5, in which almost 90% of Black children lived, were substantially lower than for White or Mixed heritage children. Rates for Asian children in quintiles 4 and 5 were between 2.5 and 6 times lower than for White children. Evidence of a stronger relationship between deprivation and rates of intervention for White than Black or Asian children was seen, echoing findings from the USA. The inverse intervention law was confirmed for White but not for Black and Asian children, perhaps because of small numbers.
4. Controlling for age, deprivation and ethnicity increased the inequality in rates between White, Asian and Black children in neighbourhoods of high deprivation. The greatest inequalities were found for children under 5 in quintile 5.

Probably the most striking findings concern the interplay of deprivation, ethnicity and age. They suggest that the question, 'Why are Black children over-represented amongst LAC?' should be replaced with two others: 'Why are Black children so concentrated in the most deprived neighbourhoods?' and 'Why are most Black children substantially *under*-represented amongst LAC (CIN and CPP) compared to White and Mixed heritage children living in similarly deprived neighbourhoods?' These findings provide a new marker of ethnic inequalities in England: overall Black children remain over-represented in the state's most powerful interventions in family life, with at best uncertain outcomes, because they are so over-represented amongst children living in areas of the highest deprivation. Racialised inequalities in the distribution, extent and consequences of deprivation must come into focus for child welfare services and more broadly.

The gap in rates between White (Black and Mixed heritage) children and Asian children was already known but is now seen to be much greater in neighbourhoods where most Asian children live. Combining CPP and LAC rates shows that in quintile 5, 1 in 50 White children were receiving a CPP or LAC intervention at 31.3.12, 1 in 88 Black children and 1 in 218 Asian children. To reiterate, the intersection between the distribution of deprivation amongst children from different ethnic groups and patterns of statutory child welfare interventions has to be the subject of further study but also new policy and practice.

A fundamental question is whether higher rates or lower rates represent greater wellbeing for children (Tilbury and Thoburn, 2009; Bywaters et al., 2014). Crudely: are Black and Asian children

under-represented amongst children on CPPs and LAC because their needs are being neglected by service providers or because there are lower levels of maltreatment? Is there something about the way Asian and Black families or communities bring up their children which results in relatively few interventions in areas of high deprivation? If so, what is there to be learnt? Or are services failing to reach children with negative short and long term consequences for their wellbeing and development? The study of health inequalities suggests that the degree of social inequality may be an explanatory factor beyond the effects of deprivation (Fone et al., 2007). Might the lower concentration of White children in deprived neighbourhoods (greater social inequality) be significant?

In any case, when such high proportions of Black and Asian children are living in the most disadvantaged neighbourhoods, questions are raised about the nature of society and the role that social workers can play in reducing or reinforcing social inequalities. When rates of extremely expensive but sometimes lifesaving interventions differ by a factor of 6 between ethnic groups but the reasons are unknown, urgent ethical, research, policy and practice issues are raised about the child welfare system.

Notes

1. Rates calculated from a NSPCC table of children on child protection plans or registers (http://www.nspcc.org.uk/Inform/research/statistics/unitedkingdom_wdf81294.pdf) and population data from the UK Census, 2011.

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Table 1: Percentage of child population living in each quintile of neighbourhoods (MSOAs) by deprivation.

Deprivation Quintile	1	2	3	4	5
Child population in England	19.4%	18.3%	18.5%	20.1%	23.7%
Distribution of MSOAs, Midlands sample	11%	20%	19%	21%	28%
Child Population Midlands Sample	12.1%	15.9%	15.7%	18.1%	38.2%
Aged 0-4	10.1%	14.1%	15.1%	18.7%	42.0%
Birmingham	2.3%	3.8%	13.0%	16.2%	64.7%
Warwickshire	30.3%	24.4%	20.5%	16.8%	8.0%

1 = most affluent 20% of neighbourhoods; 5 = most disadvantaged.

Table 2: Percentage of CIN, CPP and LAC, recorded as disabled, by age group and comparison with Blackburn et al. (2010)

Age	CIN	CPP	LAC	Blackburn et al. estimates for UK
0 to 4	3.6	0.9	1.8	3.7
5 to 9	12.3	2.8	6.3	8.2 (age 5 – 11)
10 to 15	19.4	4.0	12.7	9.5 (age 12 – 15)
16 to 17	22.6	6.0	11.2	8.5 (age 16 – 18)
Number disabled	3759	105	592	
% disabled	13.1	2.3	8.2	

Table 3: Percentage of CIN who are disabled, by age group and deprivation quintile, Midlands sample.

	Deprivation Quintile					
Age Group	1	2	3	4	5	All
0 to 4	7.1	5.3	4.5	3.3	3.0	3.6
5 to 9	25.6	14.4	12.9	13.2	10.0	12.3
10 to 15	29.5	24.2	21.5	19.9	16.5	19.4
16 to 17	30.3	29.7	23.5	20.5	20.2	22.6
All	23.5	17.6	14.9	13.2	10.7	13.1

Table 4: CIN, CPP and LAC Rates per 10,000 Children at 31.3.12 (England)

	White	Mixed	Asian	Black	Other	All
CIN Rates	306.6	559.6	197.0	493.3	318.4	315.5
CPP Rates	37.0	77.0	22.7	40.5	20.7	37.1
LAC Rates	59.2	135.3	27.6	87.3	50.4	60.4

Sources: Department for Education and ONS Annual Population Survey October 2011 – September 2012.

Table 5: CIN, CPP and LAC Rates per 10,000 Children, Midlands Sample.

	White	Mixed	Asian	Black	Other	All
CIN	253.7	351.5	109.4	226.7	298.9	235.8
CPP	39.5	62.9	21.6	34.1	37.7	37.7
LAC	64.4	122.7	17.7	71.9	51.6	60.5

Table 6: Sample Population by Ethnic Group and Deprivation Quintile - Percentages

Quintiles	1	2	3	4	5	ALL
White	15.1	19.7	18.2	19.3	27.8	100.0
Mixed	6.9	9.3	12.1	18.3	53.3	100.0
Asian	4.1	5.3	8.7	14.3	67.5	100.0
Black	1.5	2.9	6.3	12.7	76.5	100.0
Other	5.2	6.1	8.6	16.3	63.8	100.0
All	12.1	15.9	15.7	18.1	38.2	100.0

Table 7: CIN, CPP and LAC rates by ethnic group and deprivation quintile, Midlands sample.

	CIN	CIN	CIN		CPP	CPP	CPP		LAC	LAC	LAC
Quintile	4	5	ALL		4	5	ALL		4	5	ALL
White	295.9	428.0	253.7		42.4	76.8	39.5		75.5	122.1	64.4
Mixed	358.0	426.2	351.5		55.3	80.2	62.9		117.0	159.6	122.7
Asian	107.8	120.1	109.4		17.8	25.5	21.6		16.7	20.4	17.7
Black	261.4	223.3	226.7		29.9	35.3	34.1		50.8	78.3	71.9
Other	272.4	305.6	298.9		48.8	36.3	37.7		40.7	59.0	51.6
All	274.4	320.2	235.8		39.8	58.1	37.7		69.4	91.2	60.5

Table8: CPP rates by age and deprivation quintile, Midlands sample.

Quintiles	1	2	3	4	5	All
0-4	19.3	26.4	41.7	58.7	81.0	56.9
5to9	10.6	21.5	32.4	42.7	65.0	42.6
10to15	5.0	13.0	20.6	31.2	45.1	27.9
16to17	2.1	5.4	6.3	8.5	9.8	7.2
All Ages	9.5	17.6	27.8	39.8	58.1	37.7

Table 9: LAC rates by age and deprivation quintile, Midlands sample.

Quintiles	1	2	3	4	5	ALL
0-4	17.5	22.2	35.3	59.0	82.1	55.7
5to9	11.4	18.2	28.9	54.8	74.8	47.6
10to15	17.8	29.7	50.7	76.1	105.2	66.8
16to17	32.0	44.6	66.7	110.1	119.0	83.9
All Ages	17.9	26.7	42.7	69.4	91.2	60.5

Table 10: CPP, LAC and combined numbers and rates by age group, Midlands sample.

	CPP		LAC		Combined	
	Number	Rate	Number	Rate	Number	Rate
0-4	1931	56.9	1890	55.7	3821	112.7
5to9	1326	42.6	1482	47.6	2808	90.2
10to15	1091	27.9	2611	66.8	3702	94.7
16to17	99	7.2	1157	83.9	1256	91.1
All	4447	37.7	7140	60.5	11587	98.3

Table 11: Age Adjusted Inequalities in CPP Rates by Ethnicity, Quintile 5.

Age	0 to 4	5 to 9	10 to 15	16 to 17	ALL Q5	All quintiles
White	112.6	87.4	57.1	13.4	76.8	39.5
Mixed	109.6	97.2	50.9	5.9	80.2	62.9
Asian	27.7	29.6	25.0	6.7	25.5	21.6
Black	43.2	40.5	33.7	0.0	35.3	34.1
Other	54.4	37.2	23.2	0.0	36.3	37.7
All	81.0	65.0	45.1	9.8	58.1	37.7
Ratio: White/Asian	4.1	3.0	2.3	2.0	3.0	1.8
Ratio: White/Black	2.6	2.2	1.7	NA	2.2	1.2

Table 12: Age Adjusted Inequalities in LAC Rates by Ethnicity, Quintile 5.

Age	0 to 4	5 to 9	10 to 15	16 to 17	All Q5	ALL quintiles
White	110.0	106.9	140.0	137.5	122.1	64.4
Mixed	158.0	128.6	170.8	221.7	159.6	122.7
Asian	16.6	16.7	21.6	41.1	20.4	17.7
Black	53.4	56.4	107.7	130.9	78.3	71.9
Other	60.2	41.0	46.3	146.7	59.0	51.6
All	82.1	74.8	105.2	119.0	91.2	60.5
Ratio: White/Asian	6.6	6.4	6.5	3.3	6.0	3.6
Ratio: White/Black	2.1	1.9	1.3	1.0	1.6	0.9





